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Crosscorrelation Functions of DS Codes and Probability of Error Evaluations for a Multiple Access Channel

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Crosscorrelation Functions of DS Codes and Probability of Error Evaluations for a Multiple Access

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Abstract

We derive a computationally feasible closed form error expression for probability of error for a small multiple access DS-BPSK system operating in a Rayleigh fading environment. Two properties of the crosscorrelation functions of the DS codes help reduce the number of error probability evaluations required when one considers **all** possible phase shifts of the users' codes. Examination of the error rates for different phases of three maximal-length codes reveal that the probability of error is only weakly dependent on the phases.

I Introduction

The evaluation of the system capacity of a multiuser DS communication system is in general complicated. An approach is to find the maximum multiple-access interference for a given model. When the peak interference does not appear very often, this scheme yields pessimistic results. **A** more realistic approach that leads to the signal to noise ratio as a measure of performance is the use of mean-squared value of the multiple-access interference [11. However, what really measures the performance of a system and hence determines the system capacity is the probability of bit error, P_e . In a direct sequence multiple access system, P_e is a function of crosscorrelation of the direct sequence codes. In a simple but not accurate model, the effect of the multi-user interference on **the** output of the receiver correlator has been assumed to be Gaussian[2,4]. The numerical evaluation of P_e that does not invoke Gaussian assumption for multiple access interference is in general computationally intensive[5-7]. For system optimization purposes, it is of interest to know whether the relative phases of the codes have some influence on the probability of error. For instance in [SI, the authors optimize the phases of a set of codes with respect to the signal to noise ratio.

This study provides a computable closed form solution for the probability of bit error of an asynchronous direct sequence-binary phase shift keying (DS-BPSK) *code* division multiple access(CDMA) system operating in a Rayleigh fading channel with a few number of users. **It** examines whether the relative phases of the codes play a significant role in the performance of a multi-user system, and also provides some insights on the properties of crosscorrelation functions of the codes. In section 11 we specify our transmitter, channel, and receiver model. A closed form expression for the P_e is derived. In

Appendix **A** some properties of the cross correlation functions of DS codes are stated. In section III, P_e is evaluated for different phase shifts of the codes. Section **IV** concludes this study.

II Multiple Access System

Assume a star configuration with K users **connected** to a base station using a **DS-BPSK** system. In our analysis we consider user number one the as the reference user.

2.1 Transmission Model

The *kth* user information signal **is** represented by

$$
b_k(t) = \sum_{i=-\infty}^{\infty} b_i^k P_T(t - iT), \qquad (1)
$$

where b_i^k represents the k^{th} user data taking on values timing interval, $P_T(t)$ is a rectangular waveform of unit height over *(0,T*)and zero elsewhere. The pseudo-random spreading signatwe sequence with period *N* is given by from the set $\{\pm 1\}$ with equal probability during the *ith*

$$
a_k(t) = \sum_{j=-\infty}^{\infty} a_j^k P_{T_c}(t - jT_c), \qquad (2)
$$

where a_i^k represents the k^{th} user chip at the *jth* timing interval and $P_{T_{\epsilon}}(.)$ is a rectangular wave and T_c seconds duration. We assume that each bit b_i^k of $b_k(t)$ spans one period of $a_k(t)$ which takes $T = NT_c$ seconds, where T and T_c are the bit and the chip duration respectively. Each user has a different signature sequence. Then the transmitfed signal by the *kth* user is

$$
s_k(t) = Ab_k(t)a_k(t)\cos(\omega t + \phi_k),
$$
\n(3)

where ϕ_k is a random variable (r.v.) representing an arbitrary carrier phase.

2.2 Channel Model

The channel is assumed to be Rayleigh fading with

low pass equivalent impulse response
\n
$$
h_k(t) = \beta_k \delta(t - \tau_k) e^{j\varphi_k}
$$
, (4)

for the k^{th} user, where $\delta(.)$ is the Dirac delta function, τ_k is a r.v. uniformly distributed over *[O,T)* representing an arbitrary time delay, φ_k is a uniform r.v. representing an arbitrary phase shift due to the channel, β_k is a Rayleigh r.

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v. with second moment equal to $2\sigma_k^2$. All the channel variables are time invariant, mutually independent and independent from user to user. This could be an independent from user to user. approximate model of a cellular communication system. The received signal is given by

$$
r(t) = \sum_{k=1}^{K} A\beta_k b_k (t - \tau_k) a_k (t - \tau_k) \cos(\omega t + \theta_k) + \eta(t),
$$

(5)

$$
\theta_k = \phi_k + \phi_k + \omega \tau_k,
$$
 (6)

 ϕ_k and ϕ_k are such that θ_k is a r.v. uniformly distributed over $[0,2\pi)$, $\eta(t)$ is AWGN with double sided power spectral density (psd) of $N_0/2$.

2.3 Receiver Model

The receiver of the desired user is assumed to coherently recover the carrier phase and delay by locking onto the arriving desired signal[3,5]. In the receiver side, we have user number one's despreading signature, a correlator and a threshold detector. θ_1 and τ_1 , of the user *¹*received signal can be assumed to be zero with no loss of generality, The signal sample at the reference user's receiver integrator output can be called x' and is given by

$$
x' = \int_{0}^{T} r(t)a_1(t)\cos(\omega t)dt = \frac{AT}{2}\beta_1b_0^1 + \frac{AT}{2}z + v'
$$
 (7)
where $\frac{AT}{2}\beta_1b_0^1$ is the desired component, z is the multi-

AT

user interference term defined **as**

$$
z = \sum_{k=2}^{K} \beta_k \cos(\theta_k) \frac{1}{T} \int_0^T b_k(t - \tau_k) a_1(t) a_k(t - \tau_k) dt
$$
 (8)

$$
y' \sim N(0, \frac{N_0 T}{T}), \text{ and } N(m, \sigma^2) \text{ denotes a normal density}
$$

$$
v' \sim N(0, \frac{N_0 I}{4})
$$
, and $N(m, \sigma^2)$ denotes a normal density

with mean *m*, and variance σ^2 .

We can rewrite the integrator's output **as**

$$
x' = \sigma_1 \frac{AT}{2} x, \tag{9}
$$

where the decision variable x is given by

$$
x = \frac{\beta_1}{\sigma_1} b_0^1 + \frac{z}{\sigma_1} + \frac{v}{(\sigma_1 A T / 2)}.
$$
 (10)

In a Rayleigh fading environment, the actual received signal to noise ratio is
$$
\beta_1^2 \frac{E_b}{N_0}
$$
, where $E_b = \frac{A^2 T}{2}$ and the

average signal to noise ratio is

$$
\frac{E_{b1}}{N_0} = E(\beta_1^2) \frac{E_b}{N_0}.
$$
 (11)

It is clear that the third term on the right hand side of *(10)* is distributed as normal with zero mean **and** variance equal

to $\frac{1}{E_{b1}/N_0}$. The receiver decision is based on the following:

$$
x \geq 0
$$
 (12)

where the hypotheses **are**

$$
H_0: \t b_0^1 = -1 \n H_1: \t b_0^1 = +1
$$

the probability of error through the multi-user interference term **are** the continuous periodic crosscorrelation function[*1,2,9]* A set of crosscorrelation functions that contribute to

$$
\mathcal{R}_{ji}(\tau) \equiv \int_0^T a_i(t) a_j(t - \tau) dt,
$$
\n(13)

and the continuous odd crosscorrelation function
\n
$$
\hat{\alpha}_{ji}(\tau) = \int_{\tau}^{T} a_i(t) a_j(t - \tau) dt - \int_{0}^{\tau} a_i(t) a_j(t - \tau) dt.
$$
\n(14)

Let

$$
\mathcal{R}_{k1b}(\tau_k) \equiv \begin{cases} \mathcal{R}_{k1}(\tau_k) & \text{if } b_{-1}^k b_0^k = +1 \\ \hat{\mathcal{R}}_{k1}(\tau_k) & \text{if } b_{-1}^k b_0^k = -1 \end{cases}
$$
(15)

Rewriting the multi-user term z in terms of the crosscorrelation functions yields,

$$
z = \sum_{k=2}^{K} \frac{\mathcal{R}_{k1b}(\tau_k)}{T} \beta_k \cos(\theta_k). \tag{16}
$$

We note that $\beta_k \cos(\theta_k)$ is a Gaussian r.v. with zero mean and variance σ_k^2 . Therefore the conditional distribution of

z given
$$
\{\mathcal{R}_{k1b}(\tau_k), 2 \le k \le K\}
$$
 is $N(0, \sum_{k=2}^{K} \frac{\mathcal{R}_{k1b}^{\epsilon}(\tau_k)}{T^2} \sigma_k^2)$.
The detector makes a wrong decision if x is negative while $b_0^1 = +1$, or if x is positive and $b_0^1 = -1$. During the detection interval of b_0^1 , the other two bits of the k^{th} interferer namely b_{-1}^k and b_0^k for $k \neq 1$ can independently take on $\{\pm 1\}$. It is clear from equations (10), (12), (15), and (16) that,

$$
f_X(x|\beta_1, b_0^1, \mathcal{R}_{k1b}(\tau_k), 2 \le k \le K) = N(\frac{\beta_1}{\sigma_1} b_0^1, \sum_{k=2}^K \frac{\mathcal{R}_{k1b}^1(\tau_k)}{T^2} \frac{\sigma_k^2}{\sigma_1^2} + \frac{1}{E_{b1}/N_0})
$$
\n(17)

Let

$$
S_b = \{b_{-1}^k, b_0^k \mid 2 \le k \le K, b_{-1}^k \in \{\pm 1\}, b_0^k \in \{\pm 1\}\}, \qquad (18)
$$

\n
$$
S_{\tau} = \{\tau_k \mid 2 \le k \le K, 0 \le \tau_k < T\}. \qquad (19)
$$

The conditional probability of error is expressed by,

$$
P_{e|B_1, S_b, S_{\tau}} = \frac{1}{2} \{ P(x' < 0 | H_1) + P(x' > 0 | H_0) \}
$$

=
$$
Q \left(\frac{\beta_1 / \sigma_1}{\sqrt{\sum_{k=2}^{K} \frac{\alpha_{k1b}^2(\tau_k) \sigma_k^2}{T^2} + \frac{1}{E_{b1} / N_0}}} \right).
$$
 (20)

Above Q(.) is one minus the standard normal *cdf* .

has an exponential distribution with mean equal to 2,

because β_1 is a Rayleigh r.v. with second moment equal to $2\sigma_1^2$. Let

$$
\gamma_o \equiv \frac{1}{\sum_{k=2}^{K} \frac{\theta_{k1b}^2(\tau_k)}{T^2} \frac{\sigma_k^2}{\sigma_1^2} + \frac{1}{E_{b1}/N_0}}
$$
(21)

Averaging (20) with respect to β_1 yields,

$$
P_{elS_b, S_{\tau}} = \frac{1}{2} \left[1 - \frac{1}{\sqrt{1 + 1/\gamma_o}} \right]
$$
 (22)

Up on substituting for γ_a ,

$$
P_{elS_b} = \frac{1}{2} - \frac{1}{2} \int \int \int_{S_{\tau}} \cdots \int \frac{d\tau_K \dots d\tau_2}{\sqrt{1 + \sum_{k=2}^{K} \frac{\theta_{klb}^2(\tau_k)}{T^2} \frac{\sigma_k^2}{\sigma_1^2} + \frac{1}{Eb_1/No}}}
$$
(23)

As we see, due to $\mathcal{R}_{k}^2(\tau_k)$, the probability of bit error is controlled by the sign of the product $b_{-1}^k b_0^k$ instead of b_{-1}^k and b_0^k individually. The product $b_{-1}^k b_0^k$ gives the relative sign between the two bits of the k^{th} user, which are overlapping the bit of the desired user.

For a three users system, let us define the following.

$$
S_0 = \{b_{-1}^2, b_0^2, b_{-1}^3, b_0^3\} b_{-1}^2 = b_0^2, b_{-1}^3 = b_0^3\},
$$

\n
$$
S_1 = \{b_{-1}^2, b_0^2, b_0^3, b_0^3\} b_{-1}^2 = b_0^2, b_0^3 = b_0^3\}
$$
 (24)

$$
S_1 = \{b_{-1}^2, b_0^2, b_{-1}^3, b_0^3\} \, b_{-1}^2 = b_0^2, b_{-1}^3 = -b_0^3\},\tag{25}
$$
\n
$$
S_1 = \{b_1^2, b_2^2, b_3^3, b_3^3\} \, b_{-1}^2 = -b_0^2, b_3^3 = b_0^3\}.
$$

$$
S_2 = \{b_{-1}^2, b_0^2, b_{-1}^3, b_0^3\} \, b_{-1}^2 = -b_0^2, b_{-1}^3 = b_0^3\},\tag{26}
$$

$$
S_3 = \{b_{-1}^2, b_0^2, b_{-1}^3, b_0^3 \mid b_{-1}^2 = -b_0^2, b_{-1}^3 = -b_0^3\}.
$$
 (27) Hence,

$$
P_{elS_0} = \frac{1}{2} - \frac{1}{2} \int_0^1 \frac{d\tau_2 d\tau_3}{\sqrt{1 + \frac{\sigma_2^2}{\sigma_1^2} \frac{\alpha_{21}^2(\tau_2)}{T^2} + \frac{\sigma_3^2}{\sigma_1^2} \frac{\alpha_{31}^2(\tau_3)}{T^2} + \frac{1}{Eb_1/N_0}}}
$$
(28)

In (28), replacing
$$
\mathcal{R}_{31}(\tau_3)
$$
 ($\mathcal{R}_{21}(\tau_2)$) with $\mathcal{R}_{31}(\tau_3)$

 $(\hat{\ell}_{21}(\tau_2))$ gives $P_{e|S_1}$ ($P_{e|S_2}$). $P_{e|S_3}$ is obtained from (28) by replacing $\mathcal{R}_{31}(\tau_3)$ by $\hat{\mathcal{R}}_{31}(\tau_3)$ and $\mathcal{R}_{21}(\tau_2)$ by

 $\hat{\alpha}_{21}(\tau_2)$. Appendix A presents several properties of the crosssorrelation functions discussed in this section. **As** discussed in the next section, properties IV and **V** help to reduce the number of evaluations of the conditional error probabilities $P_{e|S_0}$, $P_{e|S_1}$, and $P_{e|S_2}$, when one considers all the phase shifts of the codes of the three users. Finally,

$$
P_e = \frac{1}{4} [P_{e|S_0} + P_{e|S_1} + P_{e|S_2} + P_{e|S_3}].
$$
 (29)

IU Probability of Error for Three Users

In a system with K users, assume each user has an m-sequence signature. For N=31 or 63 there are **only** three non reciprocal m-sequences available. Hence assume $K=3$ and $N=31$ from now on. Let σ_i^2 be equal for all *i*, which means each user experiences a similar Rayleigh fading. The integrals involved in the probability of error expressions such as (28) can be evaluated in closed form as described in [10]. Let $Pe_i(b, c, d)$ be the probability of bit error of the ith user given that the user number one is using its signature with phase *b,* user two with phase c, and user three with phase *d.* Because the crosscorrelation functions are **a** function of the sequences' phases, an optimization **of the** probability of error can be performed by adjusting the phases **of** the codes such that $Pe_{peak}(b, c, d) \equiv \max\{Pe_i(b, c, d): 1 \le i \le 3\}$ is minimized. **To** generate the signatures we are using a shift register configuration of **figure** 8-5 of [ll]. From [ll] we obtain the coefficients for the feedback tap *85,* **84,** 83, 82, **81,** 1. For user 1, the feedback tap is $45g=100101₂$, for user 2, the feedback tap is $75g=1111012$, and for user 3, the feedback tap is $67g=110111_2$. We need to find the values of *(b* c *d)* that achieve

 $Pe_{\text{max}} \equiv \max\{Pe_{\text{peak}}(b,c,d):1 \leq b,c,d \leq N\},$

and $Pe_{\text{min}} \equiv \min\{Pe_{\text{peak}}(b, c, d):1 \leq b, c, d \leq N\}$. Let the corresponding values of $(b c d)$ be $(b_M c_M d_M)$ and $(b_m c_m d_m)$, respectively. To do this we need to calculate Pe , N^3 times corresponding to all shifts of the three users. However, according to the property IV, $P_{e|S_0}$ **Eq.** (28) is invariant with respect to any phase shift of **any** of the codes. Therefore, we have to calculate it just once, instead of $N³$ times. Because of property V, we need to calculate P_{elS_1} and P_{elS_2} N^2 times each one, instead of N^3

 $3.4-3$ 72

times each one. Finally, P_{elS_3} needs to be calculated N^3 times. Garber and Pursley [8] performed a phase optimization for m-sequences of length 31 generated by shift registers with the same feedback taps that we have used, but they maximized the signal to noise ratio defined in [2]. For K=3, N=31 and $\sigma_1^2 = \sigma_2^2 = \sigma_3^2$ we have found

 b_M =28, c_M =16, d_M =16,
 Pe_{max} = 1.3476x10⁻² for E_b / N_o =30dB, $Pe_{\text{max}} = 3.4818 \times 10^{-2}$ for $E_b / N_o = 10 \text{dB}$; $b_m=6,~c_m=27,~d_m=12,$ $Pe_{\text{min}}^{\text{max}}$ = 3.0623x10⁻² for E_b / N_o = 10dB. $P_m=0$, $C_m=27$, $a_m=15$,
 $P_e = 0.8685 \times 10^{-2}$ for $F_b / N_s = 30$ dB,

Table I shows the phases of the codes that yield Pe_{min} and Pe_{max} and Table II from [8] shows the phases that yield minimum and maximum multi- user interference variance. It should be noted that Table **I** corresponds to the shift register configuration of Fig, 8.5 of [ll], with the most significant bit of initial loading occupying the left most cell of the shift register, whereas Table I1 corresponds to the configuration in [12]. It can be verified that the group of phases that produce minimum probability of bit error and minimum multi- user interference variance are not the same and a similar statement holds true for the maximum probability of error and the maximum interference variance. On the other hand, from these results we observe that Pe_{max} is less than two times **Pemi,.** Hence, for Rayleigh fading channel, the probability of error does not change significantly with the phase shifts of the users' codes. Our results seem to validate the results in [8], where the spread of the interference variance over different phases of the codes was also found to be of the same order, namely the ratio of maximum to minimum interference variances is less than two. Because of fading, change of *SNR* by a factor 2 would roughly translate to a factor of only 2 in error probability also. This is in contrast to an AWGN channel where the changes in codes' phases can change E_b / N_0 by as much as *2dB* at **an** error rate of 10-5[6]. *Also,* only a slight decrease in error probability, when E_b/N_0 is changed from lOdB to **3odB,** indicates that the multiple access system considered is essentially interference limited. With a processing gain of 31, the system is able to support only three users at 10^{-2} error rate. In [4] the author considers error rates when interfererers' signal range from one-fifth to one-hundredth of the reference user's signal. But the effect **of** code **phases was** not **considered.**

IV Conclusions

In this paper we derived a computationally feasible closed form error expression for a thee user **DS-BPSK** system operating in a Rayleigh fading environment. The results obtained indicate that the probability of error is only wealkly dependent on the phases of the **DS** codes.

Appendix A

Properties of Crosscorrelation Functions

.Below we state some properties of different crosscorrelation functions defined in section II. The properties can be proved by the application of periodicity of *a,; (t*) and change of variable of integration.

All the integrals that appear below are assumed to exist. *Also, g(*) and *f(*) are arbitrary functions of appropriate variables.

Let *n*, *m*, and *q* be integers,
$$
l = n-m
$$
,

$$
a_i(t) \equiv a_1(t + nT_c), \tag{A1}
$$

$$
a_j(t) \equiv a_k(t + mT_c), \tag{A2}
$$

$$
a_p(t) \equiv a_r(t + qT_c). \tag{A3}
$$

I

$$
\mathcal{R}_{k1}(\tau) = \int_{0}^{T} a_1(t) a_k(t - \tau) dt = \int_{c}^{T+c} a_1(t) a_k(t - \tau) dt.
$$
 (A4)

$$
\Pi
$$

$$
\mathcal{R}_{ji}(\tau) = \mathcal{R}_{k1}(\tau + lT_c)
$$
 (A5)

$$
\int_{0}^{T} g(\alpha_{ji}(\tau)) d\tau = \int_{0}^{T} g(\alpha_{k1}(\tau)) d\tau
$$
\n(A6)

0 T T T 0 0 This property states that $\int g(\mathbf{\ell}_{ii}(\tau))d\tau$ is invariant to any shifts of $a_k(t)$ and $a_1(t)$, since *m* and *n* are arbitrary. This means (A6) also equals $\int g(\mathbf{\ell}_{ki}(\tau))d\tau = \int g(\mathbf{\ell}_{j1}(\tau))d\tau$.

IV
\n
$$
\iint_{0}^{TT} f(\mathbf{R}_{ji}(\tau_j), \mathbf{R}_{pi}(\tau_p)) d\tau_j d\tau_p = \iint_{0}^{TT} f(\mathbf{R}_{k1}(\tau_k), \mathbf{R}_{r1}(\tau_r)) d\tau_k d\tau_r
$$
\n(A7)

$$
\int_{00}^{TT} \iint_{0} (\mathbf{R}_{ji}(\tau_j), \hat{\mathbf{\alpha}}_{pi}(\tau_p)) d\tau_j d\tau_p = \iint_{00}^{TT} \iint_{0} (\mathbf{\alpha}_{ki}(\tau_k), \hat{\mathbf{\alpha}}_{pi}(\tau_p)) d\tau_k d\tau_p
$$
\n(A8)

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TABLE I Initial Conditions of a 5-Stage Shift Register for Maximum and Minimum Probability of Error

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Chaterra na 3 ir malyniacingo

Initial Conditions of a 5-Stage Shift Register for Maximum and Minimum Interference Variance[8]

 $1.45\,h^{-1.5}$

3.4-5 **74**