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On SNR as a Measure of Performance for Narrowband Interference Rejection in Direct Sequence Spread Spectrum Systems

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On SNR **as a** Measure of Performance for Narrowband Interference Rejection in Direct Sequence Spread Spectrum Systems

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Abstract - **The usefulness of** *SNR* **as a figure of merit to quantify the narrowband interference rejection capability of a DS receiver is examined. The interference considered is a peaked autoregressive Gaussian process. The probability of error and** *SNR* **estimates of a Kalman, a modified Kalman, and a nonlinear filter proposed in [2] are obtained by simulation. Based on this simulation study and the available theoretical error rate analysis of transversal filters, we can conclude that** *SNR* **is a useful measure if the processing gain, PG, of the DS system is moderately large. When the PG is small, such as 7, and if thermal noise is negligible compared to the signal, the** *SNR* **is not a reliable measure of performance.**

I. Introduction

Narrowband interference in a received direct sequence (DS) signal can be reduced by filtering the received signal prior to despreading and demodulation (see [l] for review). The enhanced rejection of narrowband interference by the inclusion of a filter becomes more significant when the processing gain of the **DS** system is small. In several previous studies, the signal-to-noise ratio *(SNR)* obtained at the filter output has been used as a figure of performance of the interference rejection capability of a DS receiver [l], [2]. In general, *SNR* is used as a measure of performance when an analytical expression for the bit error rate *(BER)* is not available, and error rate evaluation by simulation is time consuming. However, performance comparison of two receivers based on *SNR* could be troublesome. Even though the error probability for each receiver is a monotonic function of its *SNR,* the monotonic curves of the two receivers may be different. Therefore, it is possible that the receiver showing a larger output *SNR* may have a larger error rate. In this letter, we study the suitability of *SNR* as **a** measure of performance of a **DS-BPSK** system employing one of the following for interference rejection: Kalman filter, a linear modification, the non-linear modification of [2] or a transversal filter. Simulations with a Kalman filter and its modifications are carried out to estimate the *SNR's* at the filters' output, which are sequences at the chip rate, *SNR's* of

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the test statistics, and *BER's* of the receivers. To evaluate the transversal filter, the theoretical error rate and *SNR* expressions obtained in [3] are used.

The received **DS-BPSK** signal, after chip matched filtering and sampling at the chip rate of the PN sequence, gives the following samples [2] for $k \in \{0, \pm 1, \pm 2, ...\}$, $z_k = s_k + n_k + i_k$ (1) Here $s_k = d_{k+1}c_k$, $\{d_{k+1}\}\$ denotes the data bit sequence, *L* is the PG, $\left[\frac{k}{L}\right]$ denotes the integer portion of $\frac{k}{L}$ for $k \ge 0$ and integer portion of $\frac{k}{L}$ minus 1 for $k < 0$, and $c_k \in \{-1,1\}$ is the *kth* chip of the PN sequence. For later convenience, bit d_0 is denoted as *d*. Each bit is *i.i.d* with value +1 or -1. ${n_k}$ is an i.i.d. Gaussian sequence with zero mean and variance σ_n^2 . The narrowband interference, $\{i_k\}$, is a *k L*

$$
i_k = \sum_{i=1}^p \phi_i \ i_{k-i} + e_k \tag{2}
$$

Gaussian autoregressive process of order *p* and variance σ_i^2 .

where $\{e_k\}$ is zero mean white Gaussian noise and ϕ_i 's are known to the receiver. The sequences $\{s_k\}$, $\{n_k\}$ and $\{i_k\}$ are mutually independent. In section 11, the interference filters that are simulated, are introduced along with a discussion of the key simulation variables. Section III provides a discussion and conclusions derived from this study.

11. Filtering for Narrowband Interference Rejection

The Kalman filter and its nonlinear modification are described in [2]. By putting the autoregressive interference in a state-space representation, the Kalman filter algorithm is obtained. Asymptotically, **as** the observation interval becomes large, a Kalman filter is equivalent to a Wiener filter. However, consider the situation where a decision on a given bit is made using the observations corresponding to all the past and present bits. Whereas an observation interval longer than the bit interval is useful because of the correlated interference, by extending the observation interval beyond the current bit, we also inherit the uncertainty associated with the previous bits **[3].** Hence, for certain ranges of parameters, the Kalman filter may perform poorly **as** compared to a transversal filter that has only a finite memory. Also, if the PN sequence is nearly *i.i.d*, then the

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minimum mean square error filter that estimates the interference from the observations $\{z_k\}$ is not linear [2].

A different linear filter is obtained by modifying the state space model. The received signal (1) with i_k as in (2) can be represented **as**

$$
x_{k} = \Phi_{k} x_{k-1} + w_{k}
$$
\n
$$
z_{k} = H_{k} x_{k} + n_{k}
$$
\nwhere $x_{k} = \begin{bmatrix} i_{k} & i_{k-1} & \cdots & i_{k-p+1} & d_{\lfloor \frac{k}{2} \rfloor} \end{bmatrix}^{T}$,
\n $H_{k} = \begin{bmatrix} 1 & 0 & \cdots & 0 & c_{k} \end{bmatrix}$, $w_{k} = \begin{bmatrix} e_{k} & 0 & \cdots & 0 & u_{k} \end{bmatrix}^{T}$,
\n $\Phi_{k} = \begin{bmatrix} \Phi_{1} & \Phi_{2} & \cdots & \Phi_{p-1} & \Phi_{p} & 0 \\ 1 & 0 & \cdots & 0 & 0 & 0 \\ 0 & 1 & \cdots & 0 & 0 & 0 \\ \cdots & \cdots & \cdots & \cdots & \cdots & \cdots \\ 0 & 0 & \cdots & 1 & 0 & 0 \\ 0 & 0 & \cdots & 0 & 0 & j_{k} \end{bmatrix}$

For $k = 0, \pm L, \pm 2L, \ldots, j_k = 0, u_k = \pm 1$, with equal probability, otherwise, $j_k = 1$, $u_k = 0$.

The rationale for the above modified linear filter is **as** follows. If (3) is compared with the state model of a Kalman filter in [2], it can be seen that the x_k vector has an added component $d_{\lceil \frac{k}{L} \rceil}$ as shown in (3). This reflects the knowledge that the data bit does not change over the *L* chip intervals within a data bit. This knowledge is obtained at the expense of the unknown created in the very first chip of a bit. That is, for d_0 , $j_0 = 0$ is assumed, since the true bit (which is either +1 or -1) is unknown, and $j_k = 1$ represents the fact that the bit contribution is the same for $1 \le k \le L-1$. The filtering and update equations in [2], with H , Q and Φ replaced by H_k , Q_k and Φ_k respectively, are the filtering equations for this modified filter:

$$
Q_k = \begin{bmatrix} E\{e_k^2\} & 0 & \cdots & 0 \\ 0 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \\ 0 & 0 & \cdots & u_k^2 \end{bmatrix}, \quad u_k^2 = \begin{cases} 1, k = 0, \pm L, \pm 2L, \cdots \\ 0, \text{otherwise} \end{cases}
$$

The bit decision procedure for the DS receiver is shown in Fig. **1.** The input to the PN correlator is denoted as *E,.* The output of the correlator is the test statistic *TS:*

$$
TS = \sum_{k=0}^{L-1} \varepsilon_k \, c_k \begin{array}{c} d = +1 \\ > \\ < \\ d = -1 \end{array} \tag{4}
$$

The per chip *SNR* at the output of the filter is defined as [2]: $E(e^2)$

$$
SNR_c = \frac{E(\log t)}{E(\left| \epsilon_k - s_k \right|^2)}
$$
 (5)

The test statistic *SNR* is defined **as**

$$
SNR_{TS} = \frac{E^2 \left(TS \right)}{Var(TS)}
$$
\n⁽⁶⁾

Since an interference rejection filter cannot eliminate the interference completely, ε_k has some residual correlation

Fig. 1. Filter Structure for Narrowband Interference Rejection and Bit decision of the Spread Spectrum receiver

from chip to chip, especially when the interference is strong. Hence *SNR,,* cannot be estimated from *SNR,* unless the residual correlations **and** any signal distortions are accounted for.

In a receiver with a Kalman filter, the test statistic for the current bit *d* is affected by a number of previous bits due to the filter memory. When the filter is designed for strong, highly correlated interference and thermal noise with a low variance, the intersymbol interference (ISI) introduced will be significant. Since the filter is linear, this effect may be studied by applying superposition. Let $f_K(\cdot)$ be the function of the present and all past observations defining the Kalman filter operation. The test statistic is given by

$$
TS = \sum_{k=0}^{N} [z_k - f_K(z_k)]c_k
$$
\n
$$
= \sum_{k=0}^{L-1} \Big[\{dc_k - f_K(d_{[\frac{k}{L}]}c_k)\} + \{n_k - f_K(n_k)\} + \{i_k - f_K(i_k)\} \Big]c_k
$$
\n(7)

Define $t_s = \sum_{k=0}^{L-1} \left[dc_k - f_K(d_{\left[\frac{k}{L}\right]}c_k) \right] c_k$ as the contribution of the circle component to the requirement of statistic.

the signal component to the receiver test statistic. The density of t_s provides a measure of the ISI effect on the test statistic. For example, if no filter is used, the density of t_s will be two impulses at *+L* and *-L,* indicating that the signal component of the test statistic is the true bit, i.e. no signal distortion or IS1 exists in this case. With the Kalman filter, a large variance of t_s indicates that the interference from previous bits significantly affects the current test statistic. If *t,* is Gaussian, then so is *TS.* **As** seen from the simulation results below, for certain ranges of parameters, the approximate Gaussianity of *TS* holds, whereas in certain other ranges, *TS* is either non-Gaussian or approximately Gaussian conditioned on the previous bit.

The test statistic corresponding to a nonlinear filter is, in general, non-Gaussian and its distribution is required in order to fmd the *BER.* .Any inference on the comparative performances of the linear and non-linear filters based on a test statistic *SNR* could be misleading, since the nature of the tail of the test statistics' densities of linear and nonlinear filters may be different.

A. Simulation

Simulation of the receiver in Fig. *1* is done using samples generated according to (1). Denoting *N* as the length of the PN sequence, the three cases considered are (i) $L=N=7$, (ii) *L=7, N=1023* and (iii) *L=N=63.* Case (i) is an example of small PG situations where an entire PN sequence is embedded in each bit. In case (ii), *L* is still small, but a large *N* is used so that the PN sequence is approximately *iid.* This is the situation where the nonlinear Kalman filter is expected to be better than the Kalman filter[2]. Our *BER* simulation results given below verify this. Case (iii) is for a moderate PG situation. *L* exceeding *63* was not considered because of excessive simulation time requirements. Values of σ_n^2 and σ_i^2 used are those typically expected. For example, with *L*=7, σ_n^2 =1, no filter and no interference, the *SNR* is $10\log(\frac{7^2}{7}) = 8.45dB$ (without a filter and interference, the test statistic is Gaussian and $P_b = Q(\sqrt{SNR})$, where $Q(.)$ is one

minus the standard normal cdf). For $\sigma_n^2 = 0.01$, *SNR* = *28.45dB-a* situation expected when the TX-RX pair are close. σ_i^2 =10,000 for *L*=7 and σ_i^2 =100,000 for *L*=63 correspond to a strong jammer.

The interference obeys (2) with $\phi_1 = 1.98$, $\phi_2 = -0.9801$. The spectrum of this interfering signal is sharply peaked. Sufficient number of realizations of *TS* are used to ensure that the variance in the *BER* estimate is small. We have typically used a number of samples exceeding 80/BER. More details on the simulation procedure can be found in [4].

The simulation results and some related calculations are shown in Table **I** *(L=N=7),* Table **I1** *(L=7,N=1023)* and Table **III** (*L*=N=63) for various values of σ_n^2 and σ_i^2 . The labeling of the estimates in the tables are explained below. *SNR** is the estimate of the test statistic *SNR*, *SNR_{TS}* of (6). P_b^* is the *BER* estimate obtained directly from simulations. P_{SNR}^* is an estimate of the *BER* obtained as $Q(\sqrt{SNR^*})$, i.e. under the assumption that the test statistic is Gaussian. *SNR2,* appearing in Table **I** only, is an estimate of the test statistic *SNR* obtained as *L* times the estimate of *SNR, (5),* i.e. an estimate obtained by ignoring the residual chip correlation and signal distortion. P_2^* is an estimate of the *BER* obtained via $Q(\sqrt{SNR_2})$. In the case of Kalman-based filters, the two columns correspond to filter estimates and predictor estimates respectively. In the case of transversal filters, SNR^* , P_{SNR}^* , and P_b^* shown in Table I are obtained analytically *[3].* For the case of Kalman filter, the estimated conditional densities of t_s given present bit equal to +1 and -1 are shown in Figs. *2-4.*

111. Discussion and Conclusions

A. Small Processing Gain and Short PN Sequence (L=N=7, Table **I)**

For the two linear Kalman-based filters, the *BER* estimate from the test statistic *SNR*, P_{SNR}^{*} , agrees with the *BER* estimate P_b^* for weak interference and relatively large thermal noise variance $(\sigma_n^2 = 1 \text{ and } \sigma_n^2 = 1000)$. In this situation, the **IS1** effect is small and the Gaussian assumption for the test statistic is valid. This is somewhat justified **by** the density estimate of t_s for the Kalman filter(Fig. 2). When the interference is strong and the thermal noise variance is small $(\sigma_n^2 = 0.01 \text{ and } \sigma_i^2 = 10,000)$, the variance of *t,* is large and the two peaks for each case of the true bit in the density estimate clearly show the **IS1** effect (Fig. **3). As** observed from the simulation data, the two peaks are due to the previous bit,*l*, being 1 or -1. The error estimate obtained from the conditional Gaussian approximation, $\sum \frac{1}{2} Q(\sqrt{SNR_{TS}} | l, d = +1)$, agrees with the **k{-l,l}**

BER estimate P_h^* . In the case of a transversal filter, the correct *BER* is in fact the average of the conditional errors[3]. *As* shown in Table **I** for a one-sided, 4-tap transversal filter, weak thermal noise $(\sigma_n^2 = 0.01$ and 0.1), and strong interference (σ_i^2 =10,000), the *BER* (P_h^*) is larger than the estimate $P_{SNR}^* = Q(\sqrt{SNR^*})$, where *SNR*^{*} is the average of the two conditional *SNRs.* **This** indicates significant **IS1** for these parameters. For the two linear Kalman filters, the effect of ignoring residual correlation and signal distortion can be studied by comparing the two estimates *SNR2* and *SNR*.* Irrespective of the strengths of the interference and thermal noise, the two estimates differ. Therefore, in general $SNR^* \neq L SNR_c$.

When the filter is nonlinear, the test statistic is in general non-Gaussian. However, when the interference is not strong and the thermal noise variance is relatively high, $(\sigma_n^2 = 1$ and σ^2 =1000), the Gaussian approximation to the test statistic of the nonlinear filter also yields the correct *BER* (compare P_h^*) and P_{SNR}^*). Although *SNR*^{*} of the nonlinear filter is much larger than thatof themodified linear filter or transversal filter, when $\sigma_n^2 = 0.01$ and $\sigma_i^2 = 10000$, P_b^* of the nonlinear

Fig. 2. Estimate of the conditional densities of t_s for Kalman Filter

SNR and BER Estimates											
$L = N = 7$ σ_n^2		KALMAN		MODIFIED		NONLINEAR		Trans. Filter			
σ_i^2		FILTER	PRED.	FILTER	PRED.	FILTER	PRED.	1-sided, 4 taps			
1000	$S\!N\!R^*$	5.4290	5.4290	5.1442	5.1442	5.3215	5.4254	4.7891			
	P_{SNR}^*	9.9E-3	9.9E-3	$1.2E-2$	$1.2E-2$	$1.0E-2$	9.9E-3	1.432E-02			
	SNR_2	7.3093	4.3601	7.7073	4.23	7.7429	4.411				
	P_2^*	$3.4E - 3$	$1.8E-2$	$2.7E-3$	1.9E-2	$3.2E-3$	1.8E-2				
	P_b^*	9.8E-3	9.8E-3	$1.2E-2$	$1.2E-2$	$9.5E-3$	9.5E-3	2.386E-02			
10000	SNR [*]	8.6684	8.6684	41.725	41.725	843.36	87.309	89.235			
		$1.6E-3$	$1.6E-3$	5.E-11	5.E-11	1E-185	4.E-21	1.753E-21			
	SNR_2	13.614	7.2408	8.6836	2.7871	2073.4	85.0				
	P_2^*	$1.1E-4$	$3.5E-3$	$1.6E-3$	4.7E-2		$1.E-20$				
	P_b^*	$2.7E-4$	2.7E-4	#	#	$4.0E - 5$	$4.0E - 5$	2.273E-17			
10000	SNR^*	6.8427	6.8427	6.6322	6.6322	7.9695	8.9847	20,1482			
		$4.4E-3$	4.4E-3	5.0E-3	5.0E-3	$2.3E-3$	$1.3E-3$	3.583E-06			
		12.68	6.4178	10.51	4.503	23.627	9.91				
	P_2^*	1.8E-4	$5.6E-3$	$5.9E-4$	$1.7E-2$	5.8E-7	8.2E-4				
	P_b^*	1.7E-3	$1.7E-3$	8.2E-4	8.2E-4	8.9E-3	$6.4E-3$	3.791E-04			
10000	SNR [*]	3.3297	3.3297	2.8303	2.8303	3.2994	3.4016	3.0504			
	P_{SNR}^*	$3.4E - 2$	3.4E-2	$4.6E - 2$	$4.6E - 2$	$3.4E - 2$	$3.2E-2$	4.0375E-02			
		7.007	2.954	7.481	2.831	7.1305	2.9890				
	P_2^*	$4.0E-3$	$4.3E-2$	$3.1E-3$	$4.6E - 2$	$3.8E - 3$	$4.2E-2$				
	P_b^*	3.9E-2	3.9E-2	$5.0E-2$	$5.0E-2$	3.8E-2	3.8E-2	6.583E-02			
		P_{SNR}^* P_{SNR}^* $\overline{SNR_2}$ $\overline{SNR_2}$	т.		$*$		$\overline{0}^{**}$				

Table I

 $#1$ error observed in 10⁶ trials Zero up to machine precision

filter is also higher. The nonlinear predictor and the transversal filter have the same SNR^* but the latter has a P_b^* several decades below that of the former. Even a comparison of the two nonlinear filter estimates based on *SNR*^{*} could be misleading. The nonlinear predictor and filter show almost the same *BERs (Pis)* but differ in *SNRs* by almost *1OdB.* For another set of noise parameters , σ_n^2 =0.1 and σ_i^2 =10000, the

B. Small Processing Gain and Long PN Sequence (L=7, N= 1023, Table **11)**

For the Kalman filter, the BER estimates P_{SNR}^* and P_b^*

Fig. 3. Estimate of the conditional densities of t_s for Kalman Filter

Fig. **4.** Estimate of the conditional densities of *ts* for Kalman Filter

SNR and <i>BER</i> Estimates											
$L=7, N=1023$			KALMAN		MODIFIED		NONLINEAR				
σ_n^2	σ^2		FILTER	PRED.	FILTER	PRED.	FILTER	PRED.			
1.0	1000	SNR^*	4.2585	4.2585	4.18026	4.18026	4.2388	4.3052			
		$P_{\underline{SNR}}^*$	1.95E-2	1.95E-2	2.04E-2	$2.04E-2$	1.97E-2	1.89E-2			
		P_b^*	$2.32E-2$	$2.32E-2$	2.33E-2	$2.33E-2$	2.33E-2	2.30E-2			
0.01	10000	$S\!N\!R^*$	8.083	8.083	2.699	2.699	708.548	78.598			
		$P_{\underline{SNR}}^*$	$2.23E-3$	2.23E-3	$5.02E-2$	$5.02E-2$	$0***$	3.8E-19			
		P_{b}^*	$1.63E-2$	$1.63E-2$	4.82E-2	4.82E-2	$2.2E-4$	1.9E-4			
0.1	10000	SNR^*	7.0672	7.0672	4.7866	4.7866	8.951	10.692			
		P_{SNR}^*	3.92E-3	3.92E-3	$1.43E-2$	1.43E-2	1.38E-2	5.38E-4			
		P_b^*	1.80E-2	1.80E-2	2.45E-2	2.45E-2	1.35E-2	1.24E-2			
1.0	10000	SNR^*	2.9706	2.9706	2.8912	2.8912	2.909	2.988			
		P_{SNR}^*	4.24E-2	4.24E-2	4.45E-2	4.45E-2	4.40E-2	4.19E-2			
		P_b^*	4.72E-2	4.72E-2	4.82E-2	4.82E-2	4.79E-2	4.74E-2			

Table **II** *SNR* and *BER* Etimates

Table **III** *BER* Estimates

"Zero up to machine precision

disagree for weak thermal noise and strong interference $(\sigma_n^2=0.01$ and $\sigma_1^2=10,000)$. This is due to the non-Gaussian nature of t_s as shown by its density estimate (Fig. 4). The nonlinear filter has P_b^* two decades below the others. However, based on *SNR* alone, one would have anticipated a much lower error rate for the nonlinear filter (the nonlinear filter(predictor) has a $SNR*$ 100(10) times the $SNR*$ of the Kalman). For the modified linear filter, the estimates P_{SNR}^* and P_h^* agree for all parameters considered.

C. Moderate Processing Gain and PN Sequence (L=N=63, Table III)

For moderate processing gain, the simulations had to be restricted to not too small thermal noise variances in order to observe enough errors and obtain an estimate of the BER. For all parameters considered , P_{SNR}^* and P_b^* estimates agree.

In conclusion, the *SNR* can be used **as** a measure of performance of a DS system employing narrowband interference rejection filters if the PG is moderately large. When the PG is small and the thermal noise is negligible **as** compared to the signal, *SNR* is not a reliable measure.

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